ABSTRACT: Weight-specific growth rate (G) and growth performance (the fraction of maximum growth realized, $G_{pf}$) are key demographic characteristics. The ratio of RNA/DNA (RD) can provide information on both G and $G_{pf}$. Estimating G from RD in larval fish requires an adjustment for the activity of RNA at different temperatures. Based on a meta-analysis of published data, we present a general model for the relationship between G in marine fish larvae and fluorometrically derived RD and temperature (T), and suggest that this model can be used to estimate G in marine fish larvae. Several options for estimating $G_{pf}$ are also considered, including the use of a reference growth rate ($G_{ref}$). RDs of well-fed larvae appeared to be independent of water temperatures between 4 and 28°C, suggesting that any increase in growth rate with temperature was accomplished by increased activity rather than increased concentrations of RNA. However, for the best-fit meta-analysis RD–T–G model, the relationship between RD and $G_{pf}$ was temperature dependent for fish less than fully fed.

KEY WORDS: RNA/DNA ratio · Growth · Larvae · Temperature effects · Nucleic acids · Fluorometric microplate assay · Multi-species meta-analysis
among RD, temperature (T) and G (RD–T–G relationship). These calibration experiments require specialized rearing facilities, are costly, and labor intensive. Consequently, only a few calibrations are available for a very limited number of species (Table 1), thus hampering the widespread use of RD for estimating G in fish larvae.

A variety of analytical methods have been used to estimate RNA and DNA concentrations in fish larvae. Because RD values are sensitive to the choice of procedures used, any meta-analysis of RD values from different laboratories has been problematic. This issue was addressed in a recent international intercalibration effort (Caldarone et al. 2006) whereby fluorometrically derived RD estimates were standardized based on the slopes of the RNA and DNA standard curves (Berdalet et al. 2005). This standardization method opens up both the possibility of comparing fluorometrically derived RD among laboratories, and using RD–T–G relationships developed in other laboratories.

Buckley (1984), using an UV method for analysis of RD, presented a RD–T–G calibration model based on data from larvae of 8 species of temperate marine fishes (Table 1), and suggested that there was a single relationship among RD–T–G that was independent of species. Although this general model has been used for several species, the conclusion that there is a single RD–T–G relationship in larvae of temperate marine fishes has not been further tested. Moreover, no general model based on the newer fluorometric methods has been published. To examine these and other issues, we propose the following hypotheses: (1) A single RD–T–G relationship can be used for all species of marine fish larvae, (2) RD of a well-fed larva does not change with water temperature, (3) G-values resulting from a RD–T–G relationship can be used to estimate growth performance (the fraction of maximum growth realized, Gpf), and (4) RD can be used as a direct index of condition without adjustment for the effect of water temperature on the activity of RNA. Underlying these hypotheses is a more fundamental question: How is the increase in G that sometimes accompanies an increase in temperature accomplished?

This paper addresses the above and related questions by analyzing RD data from relevant literature and by reanalyzing results from laboratory and field studies after standardizing the RD values for direct comparability. The effect of larval size and other variables on the estimation of G and Gpf using bulk nucleic acid analysis is also considered.

### METHODS

A combined data set was constructed using raw values of RD, T and G obtained from the literature. Only data which met all of the following conditions were included: (1) RD values were obtained from whole larval fish, (2) nucleic acid values were obtained from 1-dye spectrofluorometric assays, and (3) an average ratio of the slopes of the RNA and DNA standard curves was available from the researchers. All RD values were standardized (sRD) according to the procedure described in Caldarone et al. (2006) using 2.4 as the reference slope-ratio value. sRD values were used in all analyses.
A subset of the combined data set was used to examine the sRD–T–G relationship. This data set (hereafter called growth model data set, Table 2) was limited to studies where instantaneous growth coefficients were available or could be calculated using the equation:

\[ G = \frac{\ln W_2 - \ln W_1}{t_2 - t_1} \]  

where \( W_2 \) and \( W_1 \) are measures of larval size (length, dry weight, protein content) at time \( t_2 \) and \( t_1 \). When necessary, specific growth in % d\(^{-1}\) (\( G_\% \)) was converted to \( G \) by dividing values by 100, adding 1, and taking the natural log. Coefficients based on dry weights or protein were used directly. A length–weight relationship of \( W = 0.0128(L_{cm})^3 \) (where \( L_{cm} \) = length in centimeters) (Kondo et al. 1976 cited by Suda & Kishida 2003) was used to convert growth-in-length data to growth-in-weight for the 2 Japanese sardine studies, and a length exponent of 3.3 (Herzka & Holt 2000) was used for the 2 red drum studies (Table 2). Growth intervals were restricted to those ranging from 2 to 7 d. For each study, an average sRD and G for each combination of food level and temperature was calculated.

A second subset of the combined data set was used to examine sRD values of well-fed fish (hereafter called well-fed fish data set, Table 2). This data set was limited based on the following considerations: Laboratory reared larvae often exhibit an initial decrease in RD from hatching, followed by an increase through time to a plateau value, which is at or above the hatch value (Richard et al. 1991, Clemmesen 1994, Westerman & Holt 1994, Caldarone 2005). For each laboratory study included in this dataset, the average sRD value from the plateau region of the highest food-level treatment at each temperature was used. In 3 studies (all warmwater species), RD values continually increased throughout the reported sampling period (Westerman & Holt 1994, Rossi-Wongtschowski et al. 2003, Holt et al. unpubl.). For those studies, the sRD value from the last sampling day was used. To ensure that the sRD

<table>
<thead>
<tr>
<th>Species</th>
<th>Common name</th>
<th>Age range of larvae (dph)</th>
<th>Nominal temperature (°C)</th>
<th>Ratio of slopes of standard curves (mDNA:mRNA)</th>
<th>No. of data points in data set</th>
<th>Growth model</th>
<th>Well-fed-fish</th>
<th>Source</th>
</tr>
</thead>
<tbody>
<tr>
<td><em>Clupea harengus</em></td>
<td>Herring</td>
<td>5–27</td>
<td>5, 8, 11</td>
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</tr>
<tr>
<td></td>
<td></td>
<td>6–24</td>
<td>5, 6, 8, 11, 12, 13, 14</td>
<td>2.20</td>
<td>13</td>
<td>0</td>
<td>Clemmesen et al. (unpubl.)</td>
<td></td>
</tr>
<tr>
<td><em>Gadus morhua</em></td>
<td>Baltic cod</td>
<td>7–17</td>
<td>8</td>
<td>2.33</td>
<td>1</td>
<td>1</td>
<td>Grenkjaer et al. (1997)</td>
<td></td>
</tr>
<tr>
<td></td>
<td>North Sea cod</td>
<td>6–16</td>
<td>8</td>
<td>2.33</td>
<td>5</td>
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<td>St. John et al. (2001)</td>
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<td>Atlantic cod</td>
<td>6–40</td>
<td>3, 6, 9</td>
<td>2.60</td>
<td>12</td>
<td>3</td>
<td>Caldarone et al. (2003)</td>
<td></td>
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<td></td>
<td>Norwegian coastal cod</td>
<td>8–48</td>
<td>10</td>
<td>2.20</td>
<td>1</td>
<td>1</td>
<td>Skajaa et al. (2003)</td>
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<td><em>Melanogrammus aeglefinus</em></td>
<td>Haddock</td>
<td>8–48</td>
<td>6, 8, 10</td>
<td>2.68</td>
<td>12</td>
<td>3</td>
<td>Caldarone (2005)</td>
<td></td>
</tr>
<tr>
<td><em>Sardinella brasiliensis</em></td>
<td>Brazilian sardine</td>
<td>4–13</td>
<td>23</td>
<td>2.25</td>
<td>1</td>
<td>1</td>
<td>Rossi-Wongtschowski et al. (2003)</td>
<td></td>
</tr>
<tr>
<td><em>Sardinops melanostictus</em></td>
<td>Japanese sardine</td>
<td>3–15</td>
<td>19</td>
<td>1.49</td>
<td>1</td>
<td>0</td>
<td>Sato et al. (1995)</td>
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<tr>
<td></td>
<td></td>
<td>3–8</td>
<td>16</td>
<td>1.75</td>
<td>1</td>
<td>1</td>
<td>Kimura et al. (2000)</td>
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<tr>
<td><em>Sciaenops ocellatus</em></td>
<td>Red drum</td>
<td>1–14</td>
<td>28</td>
<td>3.55</td>
<td>1</td>
<td>0</td>
<td>Westerman &amp; Holt (1994) (group 1)</td>
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<tr>
<td></td>
<td></td>
<td>1–14</td>
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<td>3.55</td>
<td>1</td>
<td>1</td>
<td>Westerman &amp; Holt (1994) (group 2)</td>
<td></td>
</tr>
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<td></td>
<td></td>
<td>11–26</td>
<td>23, 28</td>
<td>5.63</td>
<td>2</td>
<td>1</td>
<td>Holt et al. (unpubl.)</td>
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<tr>
<td><em>Theragra chalcogramma</em></td>
<td>Walleye pollock</td>
<td>7–16</td>
<td>6</td>
<td>2.81</td>
<td>4</td>
<td>1</td>
<td>Canino (1997)</td>
<td></td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>55</td>
<td>13</td>
<td></td>
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</table>
values were associated with well-growing fish, only data from larvae exhibiting growth rates ≥ rates predicted by Houde & Zastrow’s (1993) growth equation for marine fish were included.

The growth model data set was used to construct growth models relating sRD and T to G. We used an information-theoretic approach (Akaike’s information criterion, AIC) for selecting the ‘best-fit’ model (Wagenmakers & Farrell 2004). Plots of residuals from the model were examined for any trends. The well-fed fish data set was used to compute a Pearson product moment correlation between sRD and T. Data were analyzed using SAS software v8.02.

Additionally, in order to reanalyze results from a published larval Atlantic cod Gadus morhua and haddock Melanogrammus aeglefinus RD–T–G study (Caldarone 2005), RD values from that study were converted to sRD values using 2.4 as the reference slope-ratio value. The sRD values were then used to recompute the multiple linear regression model published in the study. The sRD data were also fit to new models containing a sRD × T interaction term.

An independent data set (data which were not used in the model calculations) consisting of measured growth rates and RD values of herring (Clupea harengus, Hauss 2008) was used to test the fit of the meta-analysis growth model determined in the present study.

The data set consisted of larvae reared at 2 temperatures (7°C, n = 302; and 13°C, n = 147) and sampled twice: initially, and after 50 degree days. Growth rates were calculated based on dry weight and Eq. (1). Nucleic acids were analyzed with a 1-dye fluorometric assay (modified after Clemmesen 1993 and Belchier et al. 2004) with the ratio of the slopes of the standards being 2.2. RD values were standardized as outlined above.

RESULTS

Eleven published studies and 2 manuscripts in preparation met the criteria for inclusion in the combined data set (Table 2). Two subsets of this data were

<table>
<thead>
<tr>
<th>Equation no.</th>
<th>Parameters</th>
<th>sRD × T</th>
<th>Intercept</th>
<th>r²</th>
<th>F</th>
<th>K</th>
<th>ΔAIC</th>
<th>wi(AIC)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Multi-species growth-model data set, n = 61</td>
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<td></td>
<td></td>
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<tr>
<td>1</td>
<td>0.0145 ± 0.0054</td>
<td>0.0044 ± 0.0003</td>
<td>−0.0780 ± 0.0123</td>
<td>0.818</td>
<td>130</td>
<td>3</td>
<td>−417</td>
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<tr>
<td>2</td>
<td>−0.0027 ± 0.0012</td>
<td>0.0056 ± 0.0005</td>
<td>−0.0449 ± 0.0082</td>
<td>0.811</td>
<td>125</td>
<td>3</td>
<td>−415</td>
<td>2</td>
</tr>
<tr>
<td>3</td>
<td>0.0168 ± 0.0120</td>
<td>0.0006 ± 0.0023</td>
<td>−0.0836 ± 0.0289</td>
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<td>85</td>
<td>4</td>
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<td>4</td>
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<td>−0.1715 ± 0.0169</td>
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<td>6</td>
<td>0.0080 ± 0.0013</td>
<td></td>
<td>−0.0311 ± 0.0143</td>
<td>0.409</td>
<td>41</td>
<td>2</td>
<td>−348</td>
<td>70</td>
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<tr>
<td>7</td>
<td>0.0424 ± 0.0103</td>
<td></td>
<td>−0.0485 ± 0.0248</td>
<td>0.222</td>
<td>17</td>
<td>2</td>
<td>−331</td>
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<tr>
<td>Combined cod and haddock data set, n = 172</td>
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<td></td>
<td></td>
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</tr>
<tr>
<td>8</td>
<td>0.0254 ± 0.0059</td>
<td>0.0037 ± 0.0005</td>
<td>−0.0873 ± 0.0136</td>
<td>0.479</td>
<td>78</td>
<td>3</td>
<td>−1135</td>
<td>0</td>
</tr>
<tr>
<td>9</td>
<td>0.0242 ± 0.0129</td>
<td>−0.0006 ± 0.0049</td>
<td>0.0039 ± 0.0018</td>
<td>−0.0837 ± 0.0349</td>
<td>0.479</td>
<td>52</td>
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<tr>
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<td>0.0070 ± 0.0007</td>
<td>−0.0210 ± 0.0099</td>
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<td>−1132</td>
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<tr>
<td>11</td>
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<td>0.0094 ± 0.0013</td>
<td>−0.1486 ± 0.167</td>
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<td>74</td>
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<td>−1131</td>
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<td>12</td>
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<td>0.0047 ± 0.0004</td>
<td>−0.0408 ± 0.0088</td>
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<td>125</td>
<td>2</td>
<td>−1119</td>
<td>16</td>
</tr>
<tr>
<td>13</td>
<td>0.0493 ± 0.0058</td>
<td></td>
<td>−0.0793 ± 0.0157</td>
<td>0.296</td>
<td>72</td>
<td>2</td>
<td>−1085</td>
<td>50</td>
</tr>
<tr>
<td>14</td>
<td>0.0095 ± 0.0016</td>
<td></td>
<td>−0.0184 ± 0.0123</td>
<td>0.171</td>
<td>35</td>
<td>2</td>
<td>−1057</td>
<td>78</td>
</tr>
</tbody>
</table>
used for subsequent calculations: the growth model data set and the well-fed fish data set. The growth model data set consisted of 61 data points from 13 studies with sRD values ranging from 0.7 to 3.7, and instantaneous growth rates ranging from –0.06 to 0.31 at temperatures ranging from 3 to 28°C. The 7 species in this data set included starved larvae, larvae on restricted rations, and well-fed larvae on a variety of different diets (rotifers and wild plankton). Approximately 83% of the data encompassed growth rates determined from 3 to 4 d growth intervals. A plot of the raw data (Fig. 1) indicated that the slopes of the sRD–G relationship differed between temperatures, suggesting that an interaction term (sRD × T) might be necessary in the models. Therefore, growth models containing all linear combinations of sRD, T, and sRD × T were examined (Table 3). Based on the AIC values, the best-fit model for the growth data includes the interaction term, a sRD main effect term, and an intercept (Eq. 1 in Table 3). The results of the Akaike weights indicate that given the data available, there is a 56% probability that this model is the best out of the set of models tested. Based on the evidence ratio (ratio of Akaike weights, $w_i/w_j$), this model is $\sim 3 \times$ more likely to be the best model compared to the 2 next best models. A plot of the observed growth rate versus the growth rate estimated by this model suggests that the model fits the data well over the whole range of observed growth rates, including starved larvae (Fig. 2). Analysis of residuals suggests that the relationship is applicable over both the sRD and temperature ranges studied (plots not shown). Residuals from the model plotted against species showed no clear outliers (Fig. 3).

The well-fed fish data set consisted of 13 data points from 9 studies. The 6 species in this data set were considered to be ‘well-fed’ by the authors of the studies. Although the sRD data showed considerable variability, no trend in the values with temperature (Fig. 4) (Pearson correlation coefficient $r = –0.102$; $p = 0.74$) was observed. The mean of the sRD values was 3.20 ± 0.49.

After standardizing the RD values, data from calibration experiments with Atlantic cod and haddock (Caldarone et al. 2003, Caldarone 2005) were reanalyzed and models containing all linear combinations of sRD, T, and sRD × T were constructed (Table 3). Based on
AIC, the best combined cod and haddock model included the main effect of sRD, the interaction term, and an intercept (Eq. 8 in Table 3). When growth rates calculated with the original published model (Eq. 11 in Table 3) and the AIC best-fit model (Eq. 8 in Table 3) were compared, no significant difference was found (paired t-test; n = 172, p = 0.99).

At 13°C, measured growth rates for herring from the independent data set (Hauss 2008) ranged from –0.22 to 0.35 d⁻¹. At this temperature, larval growth rates estimated with the best-fit meta-analysis growth model (Eq. 1 in Table 3) fit the raw data well (average residual –0.018, Fig. 5A). The herring reared at 7°C had a much narrower range of measured growth rates (–0.20 to 0.16 d⁻¹) and a higher percentage of negative growth rates than their 13°C counterparts (35 vs. 9%). Estimated growth rates for the 7°C larvae were larger than most of the observed values (average residual –0.038, Fig. 5B).

**DISCUSSION**

**Growth estimates**

Because methodological details can affect the estimate of RD, any raw RD data derived from differing analytical protocols must first be standardized before it can be used in a meta-analysis. The fluorometric standardization method outlined in Caldarone et al. (2006) is limited to nucleic acids determined with 1-dye spectrofluorometric procedures where either the ratios of the slopes of the standards are consistent between analysis dates within the study, or the individual runs are standardized to a consistent slope ratio. Data from 2-dye spectrofluorometric nucleic acid analyses cannot be standardized using this method and were thus not included. Data was also limited to whole larvae RD to eliminate variability due to values derived from different tissues (Houlihan et al. 1988, Fernandez 1997). Working within these constraints, we were able to obtain data from a variety of larval fish species residing in a range of temperatures and fed different rations, resulting in a range of growth rates.

Few studies that define the relationship between RD, T and G in larval fish have been conducted. To date, all of the published models have been multiple linear regressions of the form:

\[
G = m_1 \times RD + m_2 \times T + C_1
\]

where G is the growth rate expressed either as an instantaneous weight- or protein-specific growth rate or in percent (G%), RD is the whole-body ratio of RNA to DNA, T is the rearing temperature in °C, m₁ and m₂ are coefficients, and C₁ is a constant (Table 1). These calibration models were mostly derived for single cool-water species reared at temperatures they normally encounter in the wild, which usually encompassed a range <7°C. In most instances, the addition of terms for larval age, size, or feeding level to these models did not significantly decrease the residual sum of squares, indicating that the relationship among RD, T and G was independent of these variables (e.g. Buckley 1982, Caldarone et al. 2003, Caldarone 2005); i.e. while RD and G are known to follow ontogenetic trends (e.g. Buckley et al. 2006, Malzahn et al. 2007, Clemmesen et al. unpub.), the RD–T–G relationship appears stable during the larval stage.

We initially fit the data from the growth model data set to a multiple linear regression model similar in form to the previously published models (Eq. 5 in Table 3). However, the plot of our meta-analysis data (Fig. 1), which included data from a much wider range of temperatures than could be tested within a single species, indicated that an interaction term between sRD and T (sRD × T) should be tested as a possible variable.
Therefore, growth models containing all linear combinations of sRD, T, and sRD × T were examined (Table 3). The model with the lowest AIC value included the interaction term, a sRD main effect term, and an intercept (Eq. 1 in Table 3). With this model, lines of constant temperature in RD–G space meet at G = –0.078, with their slopes increasing as temperature increases. A plot of the observed growth rate versus the growth rate estimated with this equation suggests that the model fits the data well over the whole range of observed growth rates, including starved larvae (Fig. 2). Analysis of residuals suggests that the relationship is applicable over the sRD and temperature range studied; however, for a given species, the model is only applicable for temperatures within the range of tolerance of that species.

Data from a previously published RD–T–G study of larval Atlantic cod and haddock was never fit to models containing an interaction term (Caldarone 2005). When this term was tested, the best fit model based on AIC included the main effect of sRD and the interaction term (Eq. 8 in Table 3). However, over the range of temperatures used in the experiments (3 to 10°C), estimates of G from the previously published model (Eq. 11 in Table 3) were not statistically significantly different from the best-fit model containing the interaction term.

Interestingly, while to our knowledge a larval fish RD–T–G relationship with an interaction term has not been previously published, there are 2 papers reporting such a model for copepods (Saiz et al. 1998, Wagn er et al. 2001). The good fit of the proposed model (Eq. 1 in Table 3) to 7 species also demonstrates an underlying similarity among fishes. Protein synthesis and the ribosomal machinery develop very early in the history of life. Ribosomal RNA sequences are highly conserved across species (Mindell & Honeycutt 1990). It should be expected, or at least not unexpected, that a single equation can describe the relationship among sRD, T and G in a single life stage across a closely related group of species (temperate marine fish). The possibility of this general relationship was already indicated by Buckley (1984) who found that a single equation adequately described the RD (measured using an UV procedure)–T–G relationship in larvae of 8 species of temperate marine fishes reared in his laboratory. The novelty of the present study comprises the (1) inclusion of warmer water species, thus expanding the temperature range analyzed, (2) inclusion of data obtained from different laboratories, and (3) development of a model which can be used by researchers analyzing nucleic acids with newer fluorometric techniques. Given the variety of fish species included in this new model, and the overall good fit given the many possible sources of variability among the growth estimates, nucleic acid analytical procedures and laboratories; it appears that use of this best-fit model of the sRD–T–G relationship should yield reasonable estimates of larval G regardless of the temperate fish species under consideration. Ideally, the RD–T–G relationship should be determined for the species of interest; however, this approach which has been undertaken by only a few researchers (perhaps due to its cost, labor intensiveness and requirement for specialized temperature-controlled rearing facilities) may not be absolutely necessary, as the meta-analysis data shows.

An independent verification of the best-fit meta-analysis model is the similarity between growth rates estimated with the model and those observed in herring reared at 13°C from the Hauss (2008) data set (Fig. 5A). The scatter of the observed growth rates around the estimated rate is most likely due to imprecision in the measurement of larval growth rate. Ideally, growth rates should be determined by measuring the change in size (protein, dry weight) of each individual. Due to the destructive nature of this method, the initial size of a larva must be estimated from the mean of a pool of larvae. However, the mean initial size is unlikely to be an accurate estimate of the individual’s initial size, given the wide range in size of larvae at a given age. At 7°C, although the model growth estimates fit the trend in the data well, the values appear to be offset by a constant amount. More laboratory RD-growth experiments especially conducted at warmer temperatures would aid in further evaluating the best-fit meta-analysis model.

**Effect of fish size and temperature on the use of RD as a condition index**

There is a large body of literature reporting RD values of fish collected in different habitats (e.g. Canino et al. 1991, Canino 1997, Clemmesen et al. 1997, Rooker et al. 1997, Kimura et al. 2000, Chícharo et al. 2003, Catalan et al. 2006, Garcia et al. 2006, Lee et al. 2006, Malzahn et al. 2007). When fish size and water temperature between sites are similar, interpretation of results is straightforward: the group with the higher RD is in better condition since it contains more machinery (potential) for protein synthesis. Interpretation of results is more problematic when size or water temperatures differ among sites. Both field and laboratory studies with well-fed fish indicate that, after initiation of exogenous feeding, growth rate is size dependent. Studies have shown a rapid initial increase in recent growth rate with larval size, followed by a more gradual increase (Otterlei et al. 1999, Buckley et al. 2006). Since the RD–T–G relationship is applicable over the
entire larval stage, it follows that as $G$ increases with larval size, so do $RD$ values. Thus, for example, a fish in optimal condition at 25 d post hatch (dph) would be expected to have a greater $G$ and $RD$ than one in optimal condition at 5 dph. Examples of size-dependent trends of $G$ and $RD$ in haddock larvae collected from Georges Bank during our GLOBEC study are illustrated in Fig. 6. The year 1999 was selected since food was plentiful in that year and process cruises were completed in 3 consecutive months — March, April and May.

Numerous published results from single-species studies and meta-analyses (e.g. Pepin 1991, Houde & Zastrow 1993) indicate that the growth rate of temperate marine fish larvae is also temperature dependent. But, what about $RD$ values — are they also temperature dependent? Or is the increase in $G$ with temperature accomplished solely through an increase in activity rather than an increase in the amount of RNA. Available $sRD$ data for well-fed larvae (well-fed fish data set) showed considerable variability, but no trend in the $sRD$ values with temperature was noted (Fig. 4). Much of the scatter around the mean (3.20) was likely due to a number of sources in addition to any species-specific effect, including differences in food quality, rearing conditions, analytical procedures, and the age or size of the larvae, although we attempted to minimize the effects of these variables (see ‘Methods’). The tight coupling among $RD$, $T$ and $G$ during the larval stage of marine fishes suggests that over periods of a few days or less adjustments in the concentration rather than changes in the activity or efficiency of RNA plays a larger role in acclimation to feeding conditions. Moreover, the lack of a trend in $sRD$ of well-fed larvae with rearing temperature suggests that at least in well-fed fish any temperature-driven increase in growth rate is accomplished through an increase in the activity rather than an increase in the concentration of RNA.

It would follow, then, that for larvae that are not food limited, direct comparison of the $RD$ values could be made regardless of the environmental temperature. In fact, they should have the same $sRD$ value ($RD_{\text{max}}$). Based on (1) the difficulty of maintaining optimal culture conditions in the laboratory, (2) the scatter we observed in $sRD$ values (Fig. 4), and (3) comparison with field data (Fig. 6), it appears that the mean $sRD$ value from our well-fed fish data set (3.2) is likely an underestimate and does not represent the $sRD$ of larvae exhibiting unconstrained growth ($RD_{\text{max}}$). Field samples from our GLOBEC study (Buckley et al. 2006) suggest that the size-dependent $RD_{\text{max}}$ approaches a $sRD$ value of about 4.4 for both Atlantic cod and haddock (Fig. 6). This higher value may also be a result of predators in the field selectively removing slower growing individuals in a cohort (Nielsen & Munk 2004), a circumstance not encountered in laboratory studies.

If condition is defined simply as the amount of protein synthetic machinery per cell, then $sRD$ values of larvae of similar size could be compared across temperatures. This approach is supported by the idea that if the larvae were brought to a common temperature, the group with the higher $RD$ value would grow faster. However, although our data from fully fed fish suggests that $RD$ values may in fact be temperature inde-
pendent, the best-fit meta-analysis model does not support this conclusion in fish less than fully fed. This aspect is discussed more fully in the following 2 sections.

### Estimating growth performance and reference growth rates

Growth performance ($G_{pf}$), the quotient of the observed growth rate ($G$) and the growth rate achieved by larva of a given size at a given temperature and photoperiod under optimal feeding and environmental conditions ($G_{max}$), provides an objective measure of larval condition. $G$ could be estimated from RD and T, or alternatively from otolith microstructure analysis. However, the effects of temperature, larval size and other factors need to be considered when determining $G_{max}$. Having a $G_{max}$ model for each species, or a generalized $G_{max}$ model based on multiple species, would be ideal. But given the difficulty in determining this value, an alternative approach would be to compare the calculated growth rates to some standard or reference growth rate ($G_{ref}$).

The available literature provides several options for estimating $G_{ref}$. Houde & Zastrow (1993) published a multi-species model based on a compilation of published data from 80 marine and estuarine species. The relationship between the midpoint of $G$ and the midpoint of T was fit to the equation:

$$G = 0.0106 \times T - 0.0023$$

for $n = 80$, $r^2 = 0.35$, and $p < 0.0001$ (corrected equation, E.D. Houde pers. comm.). This $\sim 1\%$ d$^{-1}$ increase in weight-specific growth rate for each $1^\circ C$ increase in temperature has also been seen across species (meta-analysis) in other studies, with optimal temperatures ranging from 5 to $30^\circ C$ (Morse 1989, Pepin 1991, Houde & Zastrow 1993). For individual species within their optimal temperature range, the rate of increase in weight-specific growth rate with temperature appears to vary somewhat with species and stock, but is still generally close to $1\%$ d$^{-1}$. While Eq. (3) provides an estimate of the effect of temperature on $G$ in feeding larvae across species, $G$ estimated using this equation is $<G_{max}$ since larvae fed at reduced ration were included in the analysis. Additionally, no adjustment was made for the effect of larval size on $G$.

Folkvord (2005) presented single-species size (dry weight, DW)- and temperature-dependent growth models (STDG) based on well-fed Atlantic cod larvae reared in the laboratory. He used the model for the fastest growing stock (Norwegian coastal cod, NC) to estimate reference growth rates against which cod growth performance could be measured, as follows:

$$G_{ref} = 1.20 + 1.80 \times T - 0.078 \times T \times \ln(DW) - 0.0946 \times T \times (\ln(DW))^2 + 0.0105 \times T \times (\ln(DW))^3$$

Buckley et al. (2006) suggested that the ambient photoperiod should also be considered when estimating $G_{ref}$ since photoperiod affects growth potential by delimiting the duration of daily feeding. A reanalysis of a subset of the Georges Bank Atlantic cod and haddock data set, limited to a time period when suitable prey items appeared to be present in abundance (March through May 1999, Buckley & Durbin 2006), yields the following size (protein, Pro)-, temperature-, and photoperiod (PP)-dependent growth rate (STPDG) model:

$$G = 0.01512 \ln(Pro) + 0.00686 \times T + 0.00285 \times PP - 0.08810$$

for $n = 5262$, $r^2 = 0.636$, and all parameters significant at $p < 0.0001$. This Atlantic cod and haddock STPDG model produces estimates of $G$ similar to that of Folkvord’s Atlantic cod STDG models but includes an adjustment for photoperiod (Buckley et al. 2006).

The utility of using $G$ calculated from any of the previous 3 models (Eqs. 3 to 5) as an estimate of $G_{ref}$ across species has yet to be tested, but until more general models are available, they may prove useful for estimating reference growth rates that account for the effects of temperature, size, and photoperiod. An estimated $G_{pf}$ can then be calculated by dividing observed growth rates derived from sRD values or other means, by these $G_{ref}$.

### RD, growth performance and temperature

From a convenience point of view, a RD–T–G model for which $G_{pf}$ was independent of T would be highly desirable as it would allow direct comparison of sRD values as a condition index regardless of temperature and nutritional condition. With this goal in mind we considered several G models (not shown in Table 3) where the temperature term drops out in the $G_{pf}$ calculation. One such model is $G = m_1 \times (sRD – sRD_m) \times T$ where $sRD_m$ is the (standardized) amount of RNA required for maintenance. Our reasoning was that $sRD_m$ was independent of temperature because higher maintenance costs at higher temperatures would be balanced out by higher RNA activity. This model is mathematically equivalent to $G = m_1 \times T + m_2 \times T \times sRD$ with an intercept of zero. This model gave an $r^2 = 0.77$ for the multi-species growth model data set and an $r^2 = 0.46$ for the combined cod and haddock data set. These $r^2$ are lower than those of the best-fit models in Table 3, but the zero intercept model has one less parameter.
While AIC also did not place these models among the best-fit, further consideration of these models is warranted as additional data become available.

With the best-fit meta-analysis model (Eq. 1 in Table 3) \( G_{pf} \) becomes temperature dependent as RD departs from maximum levels (Table 4). Unless additional data indicate this RD–T–G model is independent of temperature, our results along with other published studies suggest that larval size and temperature must be considered when using RD as a condition index. Thus, with the best-fit meta-analysis model, one could not assume that 2 larvae collected at different water temperatures, with equal but less than maximum RD estimates, have the same \( G_{pf} \). Additional laboratory calibration studies, especially with warm water fishes, are needed to further resolve the shape of the RD–T–G relationship and yield a conclusive answer as to whether all sRD values can be directly compared without an adjustment for temperature.

### CONCLUSIONS

After standardizing fluorometrically derived RD values (sRD), there appears to be a general equation that describes the relationship among sRD, T, and G in temperate marine fish larvae. Within the temperature range of tolerance of a given marine fish species, the general sRD–T–G relationship can be used to estimate G of larvae.

In well-fed larvae, no trend in sRD with temperature was seen. The mean value for 7 species at temperatures between 3 and 28°C was 3.20 ± 0.49 although this value is believed to be an underestimation of RD\(_{max}\).

RD is an index of \( G_{pf} \) and nutritional condition in marine fish larvae, but larval size and water temperature must be taken into account when interpreting the data.

Additional laboratory calibration experiments especially with warm water fishes, could improve the precision of the model.

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